

第 15 章 數學附錄

一 利用 demeaned form 求 $\hat{\beta}, \hat{\gamma}$

設複迴歸模型為

$$Y = \alpha + \beta X + \gamma Z + \varepsilon \quad (15A.1)$$

$$\bar{Y} = \alpha + \beta \bar{X} + \gamma \bar{Z} + \bar{\varepsilon} \quad (15A.2)$$

由式(15A.1) , (15A.2)可得：

$$Y - \bar{Y} = \beta(X - \bar{X}) + \gamma(Z - \bar{Z}) + (\varepsilon - \bar{\varepsilon})$$

$$y = \beta x + \gamma z + \varepsilon' \quad (15A.3)$$

式(15A.3)稱為 Y 的 demeaned form 的迴歸模型

現欲利用 OLS 解式(15A.3)中的 β 與 γ 。令估計的迴歸方程式為：

$$\hat{y} = \hat{\beta}x + \hat{\gamma}z$$

求 $SSE = \sum_{i=1}^n (y - \hat{\beta}x - \hat{\gamma}z)^2$ 最小

$$\frac{\partial SSE}{\partial \hat{\beta}} = 0 \quad \sum xy = \hat{\beta} \sum x^2 + \hat{\gamma} \sum xz \quad (15A.4)$$

$$\frac{\partial SSE}{\partial \hat{\gamma}} = 0 \quad \sum zy = \hat{\beta} \sum xz + \hat{\gamma} \sum z^2 \quad (15A.5)$$

根據式(15A.4) , (15A.5)解 $\hat{\beta}, \hat{\gamma}$ 可得：

$$\hat{\beta} = \frac{\sum xy \sum z^2 - \sum xz \sum zy}{\sum x^2 \sum z^2 - (\sum xz)^2} \quad , \quad \hat{\gamma} = \frac{\sum xy \sum xz - \sum x^2 \sum zy}{\sum x^2 \sum z^2 - (\sum xz)^2}$$

二 $\hat{\alpha}, \hat{\beta}, \hat{\gamma}$ 的分配

$$V(\hat{\beta}) = \frac{\sum z^2}{\sum x^2 \sum z^2 - (\sum xz)^2} \sigma^2 \quad , \quad V(\hat{\gamma}) = \frac{\sum x^2}{\sum x^2 \sum z^2 - (\sum xz)^2} \sigma^2$$

將式(16A.3)代入 $\hat{\beta}$ 得

$$\begin{aligned}
 \hat{\beta} &= \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} [\sum z^2 \sum x(\beta x + \gamma z + \varepsilon') - \sum xz \sum z(\beta x + \gamma z + \varepsilon')] \\
 &= \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} [\beta \sum x^2 \sum z^2 + \gamma \sum xz \sum z^2 + \sum z^2 \sum x\varepsilon' - \beta(\sum xz)^2] \\
 &\quad - \gamma \sum xz \sum z^2 - \sum xz \sum z\varepsilon' \\
 &= \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} [\beta[\sum x^2 \sum z^2 - (\sum xz)^2] + \sum z^2 \sum x\varepsilon' - \sum xz \sum z\varepsilon'] \\
 &= \beta + \frac{\sum z^2 \sum x\varepsilon - \sum xz \sum z\varepsilon}{\sum x^2 \sum z^2 - (\sum xz)^2} \quad (16A.6)
 \end{aligned}$$

由(15A.6)式可知 $\hat{\beta}$ 為 ε 的線性組合， ε 為常態分配，則 $\hat{\beta}$ 亦為常態分配。

$\hat{\beta}$ 之不偏性

$$\begin{aligned}
 E(\hat{\beta}) &= \beta + E\left[\frac{\sum z^2 \sum x\varepsilon - \sum xz \sum z\varepsilon}{\sum x^2 \sum z^2 - (\sum xz)^2}\right] \\
 &= \beta + \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} E[\sum z^2 \sum x\varepsilon - \sum xz \sum z\varepsilon] \\
 &= \beta + \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} (0 + 0) = \beta \\
 V(\hat{\beta}) &= [E(\hat{\beta}) - E(\beta)]^2 = E(\hat{\beta} - \beta)^2 = E\left[\frac{\sum z^2 \sum x\varepsilon - \sum xz \sum z\varepsilon}{\sum x^2 \sum z^2 - (\sum xz)^2}\right]^2 \\
 &= \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} \\
 &\quad E[(\sum z^2)^2 (\sum x\varepsilon)^2 + (\sum xz)^2 (\sum z\varepsilon)^2 - 2\sum z^2 \sum xz \sum x\varepsilon \sum z\varepsilon] \\
 &= \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} \cdot [\sigma^2 \sum x^2 (\sum z^2)^2 + \sigma^2 (\sum xz)^2 \sum z^2 \\
 &\quad - 2\sigma^2 \sum z^2 (\sum xz)^2] \\
 &= \frac{1}{\sum x^2 \sum z^2 - (\sum xz)^2} \cdot \sigma^2 (\sum z^2) [\sum x^2 \sum z^2 - (\sum xz)^2] \\
 &= \frac{\sum z^2 \sigma^2}{\sum x^2 \sum z^2 - (\sum xz)^2}
 \end{aligned}$$

上式中：

$$\begin{aligned}
 E[(\sum z^2)^2 (\sum x\varepsilon)^2] &= (\sum z^2)^2 E[(\sum x\varepsilon)^2] \\
 &= (\sum z^2)^2 E[(x_1^2 \varepsilon_1^2 + \Lambda + x_n^2 \varepsilon_n^2)] = (\sum z^2)^2 \sigma^2 \sum x^2
 \end{aligned}$$

同理可求：

$$\hat{\alpha} \sim N(\alpha, \sigma_{\hat{\alpha}}^2), \quad \hat{\gamma} \sim N(\gamma, \sigma_{\hat{\gamma}}^2)$$

$$\sigma_{\hat{y}}^2 = \sigma^2 \left[\frac{\sum x^2}{\sum x^2 \sum z^2 - (\sum xz)^2} \right]$$

三 $\Sigma(\hat{Y} - \bar{Y})^2 = \hat{\beta}\Sigma xy + \hat{\gamma}\Sigma zy$ 公式之導出

$$\begin{aligned} \Sigma(\hat{Y} - \bar{Y})^2 &= \Sigma \hat{y}^2 = \Sigma \hat{y}(y - e) = \Sigma \hat{y}y - \Sigma \hat{y}e \\ &= \Sigma \hat{y}y = \Sigma(\hat{\beta}x + \hat{\gamma}z) \cdot y = \hat{\beta}\Sigma xy + \hat{\gamma}\Sigma zy \end{aligned}$$

$$(\text{因 } \Sigma \hat{y}e = \Sigma(\hat{\beta}x + \hat{\gamma}z)e = \hat{\beta}\Sigma xe + \hat{\gamma}\Sigma ze = 0)$$

其中 $\hat{y} = \hat{\beta}x + \hat{\gamma}z$, 說明如下:

$$\text{因 } \hat{Y} = \hat{\alpha} + \hat{\beta}X + \hat{\gamma}Z, \bar{Y} = \hat{\alpha} + \hat{\beta}\bar{X} + \hat{\gamma}\bar{Z}$$

$$\hat{Y} - \bar{Y} = \hat{y} = \hat{\beta}(X - \bar{X}) + \hat{\gamma}(Z - \bar{Z}) = \hat{\beta}x + \hat{\gamma}z$$

$\Sigma xe = 0$, $\Sigma ze = 0$ 說明如下:

$$\text{因 } y = \hat{\beta}x + \hat{\gamma}z + e, \Sigma xy = \hat{\beta}\Sigma x^2 + \hat{\gamma}\Sigma xz + \Sigma xe$$

與(15A.4)比較 $\Sigma xe = 0$, 同理可證 $\Sigma ze = 0$

四 $S_{Y|XZ}^2$ 計算公式

$$\begin{aligned} \textcircled{1} S_{Y|XZ}^2 &= \frac{1}{n-k-1} \Sigma(Y - \hat{Y})^2 = \frac{1}{n-k-1} [\Sigma Y^2 - \Sigma \hat{Y}^2] \\ &= \frac{1}{n-k-1} [\Sigma Y^2 - \Sigma \hat{Y}(Y - e)] = \frac{1}{n-k-1} [\Sigma Y^2 - \Sigma Y\hat{Y} - \Sigma \hat{Y}e] \\ &= \frac{1}{n-k-1} [\Sigma Y^2 - \Sigma Y(\hat{\alpha} + \hat{\beta} + \hat{\gamma}Z)] \quad \ominus \hat{Y}e = 0 \\ &= \frac{1}{n-k-1} [\Sigma Y^2 - \hat{\alpha}\Sigma Y - \hat{\beta}\Sigma XY - \hat{\gamma}\Sigma ZY] \end{aligned}$$

$$\begin{aligned} \textcircled{2} S_{Y|XZ}^2 &= \frac{1}{n-k-1} \Sigma(y - \hat{y})^2 = \frac{1}{n-k-1} [\Sigma y^2 - \Sigma \hat{y}^2] \\ &= \frac{1}{n-k-1} [\Sigma y^2 - \Sigma \hat{y}(y - e)] = \frac{1}{n-k-1} [\Sigma y^2 - \Sigma y\hat{y} - \Sigma \hat{y}e] \\ &= \frac{1}{n-k-1} [\Sigma y^2 - \Sigma y(\hat{\beta}x + \hat{\gamma}z)] \quad \ominus \hat{y}e = 0 \\ &= \frac{1}{n-k-1} [\Sigma y^2 - \hat{\beta}\Sigma xy - \hat{\gamma}\Sigma zy] \end{aligned}$$

五 類別資料之迴歸分析(變異數分析與迴歸分析之比較)

設 X 為兩種類別的變數, 會影響依變數 Y , 兩種類別下的 Y 資料如下:

表 15a1 A、B 兩類資料

類別	Y
A	Y_{11}, \dots, Y_{1n_1}
B	Y_{21}, \dots, Y_{2n_2}

利用上述資料可進行變異數分析，並以 $\frac{MSF}{MSE}$ 比來檢定：

$$H_0 : \mu_1 = \mu_2$$

$$H_1 : \mu_1 \neq \mu_2$$

亦可利用迴歸分析檢定 X 是否會影響 Y ，在迴歸分析中設迴歸模型為：

$$Y = \alpha + \beta X + \varepsilon$$

$X = 1$ ：A 類別， $X = 0$ ：其他(B 類別)。

利用 OLS 去估計迴歸模型，可證明：

① $\hat{\alpha} = \bar{Y}_2$ ， $\hat{\beta} = \bar{Y}_1 - \bar{Y}_2$

② 檢定 $H_0 : \beta = 0$ ， $H_0 : \mu_1 = \mu_2$
 $H_1 : \beta \neq 0$ ， $H_1 : \mu_1 \neq \mu_2$

證明

① 表 15a1 的資料可寫為 X 與 Y 的表如下：

表 15a2 X 與 Y 表

	Y	X
\bar{Y}_1	Y_{11}	1
	N	N
\bar{Y}_2	Y_{1n_1}	1
	Y_{21}	0
	N	N
	Y_{2n_2}	0

於是根據表 15a2 的資料及利用 OLS 方法可估計 $\hat{\beta}, \hat{\alpha}$ 如下：

$$\begin{aligned}\hat{\beta} &= \frac{\sum xy}{\sum x^2} = \frac{\sum (X - \bar{X})(Y - \bar{Y})}{\sum (X - \bar{X})^2} = \frac{\sum XY - n\bar{X}\bar{Y}}{\sum X^2 - n\bar{X}^2} \quad (n = n_1 + n_2) \\ &= \frac{n_1\bar{Y}_1 - (n_1 + n_2)\frac{n_1}{n_1 + n_2}\frac{n_1\bar{Y}_1 + n_2\bar{Y}_2}{n_1 + n_2}}{n_1 - (n_1 + n_2)\left(\frac{n_1}{n_1 + n_2}\right)^2} = \frac{n_1\bar{Y}_1 - n_1\frac{n_1\bar{Y}_1 + n_2\bar{Y}_2}{n_1 + n_2}}{n_1 - \frac{n_1^2}{n_1 + n_2}} \\ &= \bar{Y}_1 - \bar{Y}_2\end{aligned}$$

因 $\sum X^2 = n_1$, $\sum XY = \sum_{i=1}^{n_1} Y_{1i} = n_1\bar{Y}_1$, 故

$$\begin{aligned}\bar{X} &= \frac{\sum X}{n_1 + n_2} = \frac{n_1}{n_1 + n_2} , \quad \bar{Y} = \frac{\sum Y_1 + \sum Y_2}{n_1 + n_2} = \frac{n_1\bar{Y}_1 + n_2\bar{Y}_2}{n_1 + n_2} \\ \text{又 } \hat{\alpha} &= \bar{Y} - \hat{\beta}\bar{X} = \frac{n_1\bar{Y}_1 + n_2\bar{Y}_2}{n_1 + n_2} - (\bar{Y}_1 - \bar{Y}_2)\left(\frac{n_1}{n_1 + n_2}\right) = \bar{Y}_2\end{aligned}$$

◎證明 $H_0: \beta = 0$, 即是檢定 $H_0: \mu_1 = \mu_2$
 $H_1: \beta \neq 0$, $H_1: \mu_1 \neq \mu_2$

利用 t 統計量檢定 β 如下:

$$t = \frac{\hat{\beta}}{S_{\hat{\beta}}} = \frac{(\bar{Y}_1 - \bar{Y}_2) - 0}{S_{\hat{\beta}}} = \frac{\bar{Y}_1 - \bar{Y}_2}{\sqrt{\frac{S_{Y|X}^2}{\sum x^2}}} = \frac{\bar{Y}_1 - \bar{Y}_2}{\sqrt{S_P^2\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

因 $\hat{\beta} = \bar{Y}_1 - \bar{Y}_2$, $S_{\hat{\beta}}^2 = \frac{1}{\sum x^2} S_{Y|X}^2$ 又

$$\begin{aligned}S_{Y|X}^2 &= \frac{1}{n-2} \sum (Y - \hat{Y})^2 = \frac{1}{n_1 + n_2 - 2} [\sum Y^2 - \alpha \sum Y - \hat{\beta} \sum XY] \\ &= \frac{1}{n_1 + n_2 - 2} \left[(\sum Y_{1i}^2 + \sum Y_{2i}^2 - \bar{Y}_2(n_1\bar{Y}_1 + n_2\bar{Y}_2) - (\bar{Y}_1 - \bar{Y}_2)(n_1\bar{Y}_1) \right] \\ &= \frac{1}{n_1 + n_2 - 2} [(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2] = S_P^2\end{aligned}$$

且 $\sum x^2 = \sum X^2 - n\bar{X}^2 = \frac{n_1 n_2}{n_1 + n_2}$, $\frac{1}{\sum x^2} = \frac{n_1 + n_2}{n_1 n_2} = \frac{1}{n_1} + \frac{1}{n_2}$

利用變異數分析, F 統計量為:

$$F = \frac{\sum_{i=1}^2 \sum_{j=1}^{n_i} (\bar{Y}_i - \bar{Y})^2 / 1}{\sum_{i=1}^2 \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_i)^2 / (n_1 + n_2 - 2)} = \frac{(\bar{Y}_1 - \bar{Y}_2)^2 \frac{n_1 n_2}{n_1 + n_2}}{S_P^2} = \frac{(\bar{Y}_1 - \bar{Y}_2)^2}{S_P^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)} = t^2$$

因

$$\sum_{i=1}^2 \sum_{j=1}^{n_i} (\bar{Y}_i - \bar{Y})^2 = (\bar{Y}_1 - \bar{Y}_2)^2 \left(\frac{n_1 n_2}{n_1 + n_2}\right)$$

因此得證：利用變異數分析的 $F = \frac{MSF}{MSE}$ 檢定結果與檢定兩母體的平均數是否相同的檢定完全一致。

六 若 X 分成三個類別，利用二個虛擬變數區分三個類別如下：

$X_1 = 1 : A$ 類， $X_1 = 0 : 其他$ ； $X_2 = 1 : B$ 類， $X_2 = 0 : 其他$ 。

迴歸模型建立為：

$$Y = \alpha + \beta X_1 + \gamma X_2 + \varepsilon, \text{ 且 } E(Y|X) = \alpha + \beta X_1 + \gamma X_2$$

當 $X_1 = 1$, $E(Y|X) = \alpha + \beta$ A 類 Y 的平均數

當 $X_2 = 1$, $E(Y|X) = \alpha + \gamma$ B 類 Y 的平均數

當 $X_1 = 0, X_2 = 0$, $E(Y|X) = \alpha$ C 類 Y 的平均數

其資料可表為表 15a3，利用 OLS 估計 $\hat{\beta}, \hat{\gamma}, \hat{\alpha}$ ，可得：

$$\hat{\beta} = \bar{Y}_1 - \bar{Y}_3, \hat{\gamma} = \bar{Y}_2 - \bar{Y}_3, \hat{\alpha} = \bar{Y}_3$$

又利用變異數分析方法檢定 $H_0 : \mu_1 = \mu_2 = \mu_3$ 與利用迴歸分析方法檢定 $H_0 : \beta = 0, \gamma = 0$ 完全一致。

表 15a3 三類資料

	Y	X_1	X_2
\bar{Y}_1	Y_{11}	1	0
	\vdots	\vdots	\vdots
	Y_{1n_1}	1	0
\bar{Y}_2	Y_{21}	0	1
	\vdots	\vdots	\vdots
	Y_{2n_2}	0	1
\bar{Y}_3	Y_{31}	0	0
	\vdots	\vdots	\vdots
	Y_{3n_3}	0	0